

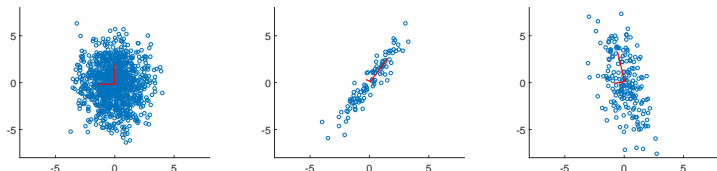
Algorithmic Trading: Functional Principal Component Analysis

Sebastian Jaimungal
University of Toronto

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PCA

- ▶ **Principal component analysis** (PCA) aims to find the **independent** modes of fluctuations of data
- ▶ For example in 2- D



- ▶ PCAs are obtained by looking at the **eigenvectors** associated with the covariance matrix of the data
- ▶ **eigenvalues** correspond to the variance in the direction of that *mode of fluctuation*

- ▶ Let Σ denote the covariance matrix of \mathbf{X}
- ▶ Since Σ is positive-definite, there exists an **orthogonal matrix** U and a **diagonal matrix** D such that

$$\Sigma = U D U^T$$

- ▶ Moreover, U corresponds to the matrix of **eigenvectors** of Σ (stacked columnwise) and the diagonal elements of D the corresponding **eigenvalues**.

$$\Sigma U_{\cdot k} = D_{kk} U_{\cdot k}$$

$$\text{and } U_{\cdot j}^T U_{\cdot k} = \delta_{jk}$$

- ▶ Hence, writing $\mathbf{Y} = \mathbf{U}^T \mathbf{X}$,

$$\begin{aligned}
 \mathbb{C}[\mathbf{Y}_i, \mathbf{Y}_j] &= \mathbb{C} \left[\sum_k \mathbf{U}_{ki} \mathbf{x}_k, \sum_l \mathbf{U}_{lj} \mathbf{x}_l \right] \\
 &= \sum_{k,l} \mathbf{U}_{ki} \mathbf{U}_{lj} \mathbb{C}[\mathbf{x}_k, \mathbf{x}_l] \\
 &= \sum_{k,l} \mathbf{U}_{ki} \mathbf{U}_{lj} \boldsymbol{\Sigma}_{kl} \\
 &= (\mathbf{U}^T \boldsymbol{\Sigma} \mathbf{U})_{ij} \\
 &= (\mathbf{U}^T \mathbf{U} \mathbf{D} \mathbf{U}^T \mathbf{U})_{ij} = \mathbf{D}_{ij}
 \end{aligned}$$

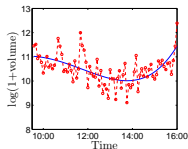
- ▶ The elements of the random variables \mathbf{Y} are all **independent**, and have variance \mathbf{D}_{ij}
- ▶ They are the **principal components**

FPCA

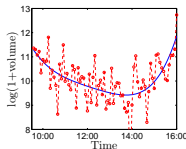
- **Functional Principal Component Analysis (FPCA)** is the functional version of PCA
- The raw data $\{x_{i1}, \dots, x_{in_i}\}_{i=1, \dots, N}$ are viewed as noisy observations of a function at discrete points in “time” $\{t_{i1}, \dots, t_{in_i}\}_{i=1, \dots, N}$
- Take observation i , and **regress** the **observed values** $\{x_{ij}\}_{j=1, \dots, n_i}$ **onto a basis** $\phi(t) = \{\phi_1(t), \dots, \phi_K(t)\}$, i.e.

$$\alpha_i = \arg \min_{\alpha} \sum_{j=1}^{n_i} \left(x_{ij} - \sum_{k=1}^K \alpha_{i,k} \phi_k(t_{ij}) \right)^2$$

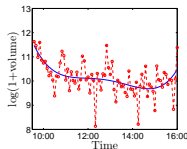
this can be done using **least-squares**.



(a) Oct-02-2014



(b) Oct-03-2014



(c) Oct-04-2014

Many different basis choice... domain specific knowledge

- **Legendre polynomials** $P_n(t)$ basis, these are solutions to Legendre's ODE

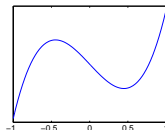
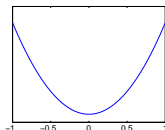
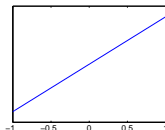
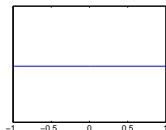
$$\frac{d}{dt} \left((1-t^2) \frac{d}{dt} P_n(t) \right) + n(n+1) P_n(t) = 0$$

and can be written as

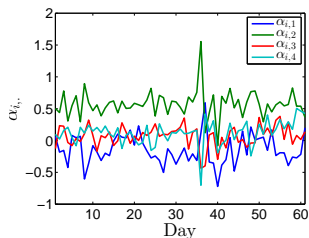
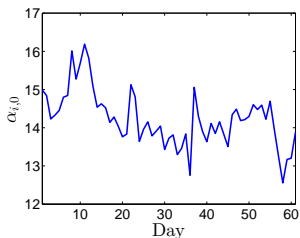
$$P_n(t) = \frac{1}{2^n n!} \frac{d^n}{dt^n} \left((t^2 - 1)^n \right)$$

They form an **orthogonal basis** w.r.t. the L^2 inner product

$$\int_{-1}^1 P_m(t) P_n(t) dt = \frac{2}{2n+1} \delta_{mn}.$$



This provides a “time-series” of coefficients $\alpha_{i,j}$.



FPCA typically assumes that these are independent.. but there are modifications that account for functional time-series.
We will stick to the independent case.

- Define the **covariance function**

$$K(s, t) = \text{cov}(x(s), x(t))$$

this is the analog of the **covariance matrix** Σ of a multi-variate sequence of observations

- Recall for PCA

$$\Sigma = \mathbf{U}^T \mathbf{D} \mathbf{U} = \sum_{j=1}^M d_j \mathbf{U}_{\cdot j} \mathbf{U}_{\cdot j}^T \quad \text{with} \quad \mathbf{U}_{\cdot i}^T \mathbf{U}_{\cdot j} = \delta_{ij}$$

- **Mercer's lemma** states that if $K(s, t)$ is continuous on \mathcal{D}^2 , then there exists an **orthonormal** sequence of continuous functions $\{\psi_i(t), i = 1, \dots\}$ (**eigenfunctions**) such that

$$(\mathcal{K} \psi_i)(t) = \kappa_i \psi_i(t),$$

where the **kernel operator** \mathcal{K} acts on square-integrable functions f as follows

$$(\mathcal{K} f)(t) = \int_{\mathcal{D}} K(s, t) f(s) ds.$$

Moreover, the covariance function

$$K(s, t) = \sum_{j=1}^{\infty} \kappa_j \psi_j(s) \psi_j(t), \quad \text{with} \quad \int_{\mathcal{D}} \psi_i(t) \psi_j(t) dt = \delta_{ij}$$

$$\text{and } \sum_{j=1}^{\infty} \kappa_j = \int_{\mathcal{D}} K(s, s) ds < +\infty.$$

- The **Karhunen-Loève** expansion states that under the assumptions of Mercer's lemma, we have

$$x(t) = \mu(t) + \sum_{j=1}^{\infty} \sqrt{\kappa_j} \xi_j \psi_j(t),$$

where,

$$\xi_j = \frac{1}{\sqrt{\kappa_j}} \int_{\mathcal{D}} x(s) \psi_j(s) ds$$

is a random variable with

$$\mathbb{E}[\xi_i] = 0 \quad \text{and} \quad \mathbb{E}[\xi_i \xi_j] = \delta_{ij}, \quad \forall i, j \in \mathbb{N},$$

and the series converges uniformly on \mathcal{D} w.r.t the L^2 norm.

- ▶ The eigenvalues κ_i can be interpreted as the size of the variation of $x(t)$ in the direction of $\psi_i(t)$
- ▶ We typically truncate the series to obtain a **dimensionally reduced approximation** of x

$$\hat{x} = \mu + \sum_{j=1}^M \sqrt{\kappa_j} \xi_k \psi_k(t)$$

- ▶ But **how to estimate** the eigenfunctions and eigenvalues from data?

- ▶ Let the **sample mean function** be denoted $\bar{\mu}$, i.e.,

$$\bar{\mu}(t) = \frac{1}{n} \sum_{i=1}^n x_i(t)$$

- ▶ The empirical approximation of $K(s, t)$ is

$$\hat{K}(s, t) = \frac{1}{N} \sum_{i=1}^N (x_i(s) - \mu(s)) (x_i(t) - \mu(t)) .$$

- ▶ Mercer's lemma then implies there exists an orthonormal basis of eigenfunctions $\hat{\psi}_j$ with eigenvalues $\hat{\kappa}_j$ such that

$$\hat{K}(s, t) = \sum_{j=1}^{\infty} \hat{\kappa}_j \hat{\psi}_j(s) \hat{\psi}_j(t)$$

- Expand the (centered) observations and eigenfunctions onto the basis functions

$$x_i(t) = \sum_{k=1}^K c_{ik} \phi_j(t), \quad \text{and} \quad \hat{\psi}_j(t) = \sum_{k=1}^K b_{jk} \phi_j(t)$$

- Then one can show that

$$\mathbf{b}_j = \mathbf{W}^{-\frac{1}{2}} \mathbf{u}_j$$

where \mathbf{u}_j solves the eigenproblem

$$\left(\frac{1}{N} \mathbf{W}^{\frac{1}{2}} \mathbf{C}^{\top} \mathbf{C} \mathbf{W}^{\frac{1}{2}} \right) \mathbf{u}_j = \kappa_j \mathbf{u}_j$$

and

$$W_{kl} = \langle \phi_k, \phi_l \rangle .$$

Figure: INTL (2014) traded volume using 5 minute buckets.

